Childhood maltreatment is associated with distinct genomic and epigenetic profiles in posttraumatic stress disorder

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Childhood maltreatment is associated with distinct genomic and epigenetic profiles in posttraumatic stress disorder


Childhood maltreatment is likely to influence fundamental biological processes and engrave long-lasting epigenetic marks, leading to adverse health outcomes in adulthood. We aimed to elucidate the impact of different early environment on disease-related genome-wide gene expression and DNA methylation in peripheral blood cells in patients with posttraumatic stress disorder (PTSD). Compared with the same trauma-exposed controls (n = 108), gene-expression profiles of PTSD patients with similar clinical symptoms and matched adult trauma exposure but different childhood adverse events (n = 32 and 29) were almost completely nonoverlapping (98%). These differences on the level of individual transcripts were paralleled by the enrichment of several distinct biological networks between the groups. Moreover, these gene-expression changes were accompanied and likely mediated by changes in DNA methylation in the same loci to a much larger proportion in the childhood abuse (69%) vs. the non-child abuse-only group (34%). This study is unique in providing genome-wide evidence of distinct biological modifications in PTSD in the presence or absence of exposure to childhood abuse. The findings that nonoverlapping biological pathways seem to be affected in the two PTSD groups and that changes in DNA methylation appear to have a much greater impact in the childhood-abuse group might reflect differences in the pathophysiology of PTSD, in dependence of exposure to childhood maltreatment. These results contribute to a better understanding of the extent of influence of differences in trauma exposure on pathophysiological processes in stress-related psychiatric disorders and may have implications for personalized medicine.

epigenome | biomarkers | psychiatry | development

Childhood maltreatment is a complex problem that exerts an enormous impact on individuals, families, and society, and is of great significance for public health. Maltreatment during childhood likely influences fundamental biological processes and engraves long-lasting epigenetic marks, leading to adverse health outcomes in adulthood (1). Exposure to adverse life events in childhood has not only been linked to an increased proportion in the childhood abuse (69%) vs. the non-child abuse-only group (34%). This study is unique in providing genome-wide evidence of distinct biological modifications in PTSD in the presence or absence of exposure to childhood abuse. The findings that nonoverlapping biological pathways seem to be affected in the two PTSD groups and that changes in DNA methylation appear to have a much greater impact in the childhood-abuse group might reflect differences in the pathophysiology of PTSD, in dependence of exposure to childhood maltreatment. These results contribute to a better understanding of the extent of influence of differences in trauma exposure on pathophysiological processes in stress-related psychiatric disorders and may have implications for personalized medicine.

Results

In the present study we selected 169 individuals, all of whom had experienced at least two types trauma other than childhood abuse. Of these individuals, 108 had a trauma history but did not have a life-time history of PTSD or current PTSD symptoms, and these were thus considered trauma-exposed controls. Sixty-one individuals met criteria for current PTSD, of which 32 reported a history of childhood maltreatment and 29 did not report childhood abuse. Clinical and epidemiological demographics of the participants across the three investigated groups are depicted in Table 1. To control for the observed significant group differences in age, sex, ethnicity, adult trauma severity, and substance abuse, these variables were used as covariates in the association analyses.

Effects of Childhood Maltreatment on PTSD-Associated Transcriptional Profiles

To test for gene-expression differences between individuals with PTSD and trauma-exposed controls, linear regression


Conflict of interest statement: K.J.R. has an unrelated role as cofounder of Extinction Pharmaceuticals for development of N-methyl-o-aspartate-based therapeutics.

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models were built to assess the influence of the case status on expression profiles. A total of 303 transcripts were differentially expressed between PTSD individuals with child abuse vs. controls, and 244 transcripts were differentially expressed between PTSD individuals without child abuse vs. controls after corrections for multiple testing (adjusted \( P < 0.05 \)) (Fig. 1 and Dataset S1). An overlap of only 14 transcripts was observed (2%), indicating distinct biomarker profiles in PTSD in the presence or absence of child abuse. This overlap is no more than expected by random chance (Fisher’s exact \( P = 0.095 \)). Post hoc analysis identified that only 4% of the differentially expressed transcripts might be confounded by lifetime PTSD treatment and only 1.5% of the transcripts might be confounded by different proportions of immune cells types, suggesting that these factors are not major confounds in the current analysis (Dataset S1). Among the differentially expressed transcripts, we successfully validated the gene-expression changes in the corresponding groups by using quantitative PCR (qPCR) in five of five tested transcripts (Dataset S2).

**Correlation of Gene-Expression Differences with DNA Methylation Changes.** Childhood maltreatment has an influence on biological processes via epigenetic modifications (22). To identify if the above gene-expression changes were correlated with DNA methylation differences in loci encoding these transcripts, we investigated DNA methylation profiles of the transcripts differentially regulated in either of the PTSD case groups versus controls (\( n = 547 \)) using the Illumina 450 k HumanMethylation array. Transcripts with at least one CpG within their respective gene locus on the Illumina HumanMethylation 450 k array (\( n = 304 \)) were assessed for methylation differences in the same subjects tested for gene-expression differences. Although at least one significantly differentially methylated CpG site was identified in 69.3% of transcripts unique to PTSD cases with child abuse, only 33.6% of transcripts unique to PTSD cases without child abuse also showed differences on the DNA methylation level for at least one CpG after correction for multiple testing. This difference was even more accentuated in transcripts for which the locus showed significant differences in DNA methylation in five or more CpGs; here, 11.7% of the transcripts showed such matching epigenetic differences in the PTSD child abuse group, but this was only the case for 0.8% of the transcripts from the PTSD non-child-abuse-only group (Fig. 2A and B, and Dataset S3). DNA methylation results were validated for seven CpG sites in two distinct loci using an independent method, the Sequenom EpiTYPER (Dataset S4).

The relation between DNA methylation and gene expression is complex. In line with other studies (23, 24), we observe negative correlations of CpGs with gene expression close to the transcription start sites and an increasing proportion of positive correlations in proximal and distal regulatory sites (Fig. 2B and Dataset S5).

Compared with all tested CpGs within the differentially expressed transcripts, the CpGs that also show significantly different methylation levels are more often located close to the 3′ UTR and in the body of the gene as well as in the shelves of the CpG islands and the open sea. In addition, the vast majority of these CpGs were hypermethylated in the case group (Fig. 2C). In addition, even though the absolute differences in methylation levels between the groups are small (for most less than 5%), the direction of the differences in DNA methylation between the case and the control groups agrees to a large extent (73% for individual CpGs and 94% at the locus level) with the observed group differences in gene expression (see Fig. 3 and Dataset S5 for more detail).

These results indicate that changes in gene expression are more often associated with changes in DNA methylation in PTSD occurring after child abuse than PTSD occurring in the absence of child abuse, and that these changes in DNA methylation may indeed mediate the observed gene-expression changes.  

**Table 1. Distribution and comparisons of demographic and clinical variables of the samples included in the study**

<table>
<thead>
<tr>
<th>Sample</th>
<th>Controls (( n = 108 ))</th>
<th>PTSD + child abuse (( n = 32 ))</th>
<th>PTSD no child abuse (( n = 29 ))</th>
<th>( P ) value</th>
<th>( P ) value</th>
<th>( P ) value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age – mean (SD)</td>
<td>44.23 (1.2)</td>
<td>39.56 (1.8)</td>
<td>43.69 (2.0)</td>
<td>0.059</td>
<td>0.044</td>
<td>0.134</td>
</tr>
<tr>
<td>Sex</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Men</td>
<td>30 (27.8%)</td>
<td>5 (15.6%)</td>
<td>13 (44.8%)</td>
<td>0.244</td>
<td>0.113</td>
<td>0.023</td>
</tr>
<tr>
<td>Women</td>
<td>78 (72.2%)</td>
<td>27 (84.4%)</td>
<td>16 (55.2%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Ethnicity</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>African American</td>
<td>101 (93.5%)</td>
<td>24 (75%)</td>
<td>25 (86.2%)</td>
<td>0.006</td>
<td>0.245</td>
<td>0.343</td>
</tr>
<tr>
<td>Other</td>
<td>7 (6.5%)</td>
<td>8 (25%)</td>
<td>4 (13.8%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PSS – mean (SD)</td>
<td>3.49 (0.32)</td>
<td>35.06 (1.2)</td>
<td>29.83 (1.7)</td>
<td>1.89 ( \times 10^{-7} )</td>
<td>1.14 ( \times 10^{-52} )</td>
<td>0.011</td>
</tr>
<tr>
<td>CAPS current – mean (SD)</td>
<td>26.19 (3.3)</td>
<td>112.73 (9.8)</td>
<td>79.71 (10.2)</td>
<td>1.68 ( \times 10^{-17} )</td>
<td>4.66 ( \times 10^{-9} )</td>
<td>0.054</td>
</tr>
<tr>
<td>PTSD diagnosis</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>108 (100%)</td>
<td>0</td>
<td>0</td>
<td>2.59 ( \times 10^{-32} )</td>
<td>1.4 ( \times 10^{-26} )</td>
<td>1</td>
</tr>
<tr>
<td>Yes</td>
<td>0</td>
<td>32 (100%)</td>
<td>29 (100%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>BDI mean (SD)</td>
<td>8.08 (0.7)</td>
<td>30.13 (2.3)</td>
<td>25.27 (2.1)</td>
<td>2.36 ( \times 10^{-24} )</td>
<td>8.47 ( \times 10^{-18} )</td>
<td>0.128</td>
</tr>
<tr>
<td>Child abuse (CTQ) – mean(SD)</td>
<td>34.35 (1.0)</td>
<td>76.23 (3.5)</td>
<td>37.24 (2.0)</td>
<td>4.69 ( \times 10^{-32} )</td>
<td>0.206</td>
<td>5.15 ( \times 10^{-9} )</td>
</tr>
<tr>
<td>No of types of child abuse</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(moderate to severe)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>None</td>
<td>75 (69.4%)</td>
<td>0</td>
<td>29 (100%)</td>
<td>1.19 ( \times 10^{-21} )</td>
<td>0.317</td>
<td>2.54 ( \times 10^{-13} )</td>
</tr>
<tr>
<td>One</td>
<td>21 (19.4%)</td>
<td>0</td>
<td>0</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Two</td>
<td>12 (11.2%)</td>
<td>32 (100%)</td>
<td>0</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Total types adult trauma</td>
<td>4.63 (0.23)</td>
<td>7.91 (0.6)</td>
<td>6.49 (0.6)</td>
<td>4.69 ( \times 10^{-9} )</td>
<td>0.005</td>
<td>0.154</td>
</tr>
<tr>
<td>(TEI) – mean (SD)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Substance abuse current</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>103 (95.4%)</td>
<td>28 (87.5%)</td>
<td>25 (86.2%)</td>
<td>0.111</td>
<td>0.077</td>
<td>0.881</td>
</tr>
<tr>
<td>Yes</td>
<td>5 (4.6%)</td>
<td>4 (12.5%)</td>
<td>4 (13.8%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PTSD treatment lifetime</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>108 (100%)</td>
<td>24 (72%)</td>
<td>24 (82.8%)</td>
<td>8.9 ( \times 10^{-5} )</td>
<td>0.03</td>
<td>0.352</td>
</tr>
<tr>
<td>Yes</td>
<td>0</td>
<td>8 (28%)</td>
<td>4 (17.2%)</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
The objective of this study was to interrogate biomarker profiles at the level of gene expression and DNA methylation in individuals with PTSD with or without childhood abuse. We observed that although robust gene-expression differences could be identified for both PTSD groups compared with trauma-exposed controls, the significant changes hardly overlapped at the transcript-specific level (2%) and only modestly (35%) at the biological pathway level. In fact, the differentially expressed transcripts in each group showed highly variable regulation and often the opposite pattern between the two groups. The use of two distinct but trauma-matched controls groups would make it more difficult to determine if the differences originate in the case or the control groups. We acknowledge that this approach has limitations and therefore performed a post hoc analysis using matched controls with or without child abuse and did not see any differences in the percentage of overlap in differentially expressed genes (Dataset S7) between the two PTSD case group comparisons, indicating that these results are not likely confounded by the study design. Additionally, comparisons of the transcripts defining the signature of child abuse [by (i) direct comparison within PTSD cases and (ii) comparison of individuals with and without child abuse using the whole sample] with transcripts significantly different between any of the two PTSD case groups and controls showed a minimal overlap (0.7%), suggesting that the gene-expression signature of PTSD with and without child abuse is distinct from the signature of child abuse within PTSD cases (Dataset S7).

Another potential confound of these results is the possibility that the observed gene-expression and DNA methylation profiles are only a reflection of differences in peripheral blood cell types. Comparison with transcripts correlated with the relative amount of the main immune cell subtypes (CD15–CD16, CD14, CD4, and CD8) revealed that only up to 1.5% of the transcripts might be confounded by differential cell counts. The possibility of the

Enrichment of Biological Processes Among Differentially Expressed Transcripts. As we observed differences in gene-expression profiles and epigenetic marks in PTSD occurring in the presence or absence of childhood maltreatment, we next sought to better comprehend how the transcripts interacted with each other. We built subnetworks from the transcript relationships using a modified gene-set enrichment analysis to identify overrepresented networks in Ariadne Pathway Studio 8.0.

Although functional annotation revealed overlapping pathways between the groups, such as cell survival, cell development, cell migration, cell adhesion, T-cell activation, and immunity networks, several nonoverlapping networks were observed between the groups. The most notable differences included the enrichment of the central nervous system development and tolerance induction pathways in the PTSD group with child abuse and the apoptosis and growth rate networks in the PTSD group without child abuse (Fig. S1 and Dataset S6).

Interestingly, despite the completely divergent transcripts differentially expressed between the two groups, several overlapping networks were identified, suggesting a possible common downstream effect of disease in blood cells.

Discussion

The analyses were performed using a common set of trauma-exposed controls, not matched by exposure to childhood maltreatment. This approach allowed identifying distinct biomarker profiles between the two case groups. The use of two distinct but trauma-matched controls groups would make it more difficult to determine if the differences originate in the case or the control groups. We acknowledge that this approach has limitations and therefore performed a post hoc analysis using matched controls with or without child abuse and did not see any differences in the percentage of overlap in differentially expressed genes (Dataset S7) between the two PTSD case group comparisons, indicating that these results are not likely confounded by the study design. Furthermore, comparisons of the transcripts defining the signature of child abuse [by (i) direct comparison within PTSD cases and (ii) comparison of individuals with and without child abuse using the whole sample] with transcripts significantly different between any of the two PTSD case groups and controls showed a minimal overlap (0.7%), suggesting that the gene-expression signature of PTSD with and without child abuse is distinct from the signature of child abuse within PTSD cases (Dataset S7).

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Participants in this study belonged to a larger study investigating the impact of exposure to type of trauma. This finding could not only have implications for biomarker research for stress-related disorders but may help to elucidate pathophysiological differences in dependence of trauma exposure in immune disturbances often accompanying these disorders (3, 4, 30, 31). If these distinct changes are not limited to peripheral blood cells but extend to other organ systems, as indicated by some studies (27, 28, 32), they might also aid in the search for disease mechanisms and therapeutic intervention.

Materials and Methods

Samples. Participants in this study belonged to a larger study investigating the contribution of genetic and environmental factors in PTSD (33). All study procedures were approved by the institutional review boards of Emory University School of Medicine and Grady Memorial Hospital and all subjects gave written informed consent to the study. The 396 participants were selected from a larger study, which has been described previously (34). Microarray gene-expression results had been reported in a previous report (34) for a subset of these patients (n = 211). For this study, we selected patients from a now expanded set of microarray gene-expression results (adding 185 new samples with gene-expression array data for all 396 individuals).

To the best of our knowledge, this genome-wide study assessing the influence of exposure to childhood maltreatment on gene expression and DNA methylation profiles in peripheral blood in PTSD is unique. Our findings are in line with a recent study demonstrating that epigenetic alterations in the hippocampus among suicide completers with or without childhood trauma were distinct (29), although this study did not explore related changes in gene expression. These data provide further evidence that there are substantial differences in disease-related biological perturbations in the presence or absence of exposure to maltreatment in childhood. The fact that several nonoverlapping biological pathways seem to be affected in the two PTSD groups and that changes in DNA methylation appear to have a much greater impact in the childhood maltreatment group may indicate differences in the pathophysiology of PTSD, in dependence of biological pathways seem to be affected in the two PTSD groups and that changes in DNA methylation appear to have a much greater impact in the childhood maltreatment group may indicate differences in the pathophysiology of PTSD, in dependence of

impact of other specific immune-cell subtypes, not measured in our analysis can, however, not be excluded. Furthermore, because the gene-expression levels and methylation patterns were only measured at one time-point, it is difficult conclude whether the expression and methylation changes are indeed the direct consequence of exposure to child abuse per se or a reflection of other factors changed by the response to child abuse or related to the disease itself.

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Psychological Assessments. PTSD symptomatic scale and clinician administered PTSD scale. The modified PTSD symptomatic scale (mPSS) was used as a measure of PTSD symptoms (35). PSS frequency items were summed to obtain a continuous measure, with values of above 20 considered as clinically significant PTSD symptoms. In addition, a categorical PTSD classification was made by using the scores of the B, C, and D clusters of the PSS. Individuals classified as having PTSD needed to have clinically significant symptoms in all three clusters, representing intrusive, avoidance/numbing, and hyperarousal (36). The clinician-administered PTSD scale (CAPS) was available for a subset of 349 of the 396 individuals (37). There was a significant strong positive correlation between PSS and current CAPS in this dataset (Pearson $r = 0.69, P = 3.15 \times 10^{-29}$). The CAPS was also used to assess the presence or absence of lifetime PTSD.

Beck depression inventory. The Beck depression inventory (BDI) was administered to measure current depression symptoms (38). Current depression and PTSD symptoms were highly correlated ($r = 0.653, P = 1.56 \times 10^{-46}$). Diagnostic tests for variance inflation factor and tolerance revealed a variance inflation factor of 1.72 and tolerance of 0.58, indicating no evidence for multicollinearity between depressive and PTSD symptoms.

Substance abuse. As described previously (34), two self-report items were used to assess for current alcohol and substance use-related problems.

Trauma-Exposure Measurements. Childhood trauma questionnaire. The childhood trauma questionnaire (CTQ) was used as a measure of child abuse. The CTQ is a consistent and stable self-report inventory assessing five types of childhood trauma: sexual, physical, and emotional abuse, and emotional and physical neglect (39). In accordance with prior research, we restricted our analyses to the three abuse subscales only (40). As described previously, participants were dichotomized into two groups for each of the three categories of abuse (presence or absence of moderate to severe physical, sexual, or emotional abuse). Finally, we created a composite variable across the three abuse types, grouping participants into (i) those with no exposure to any childhood abuse (scores less than cut-off for all three abuses), and (ii) those with moderate to severe exposure to at least two types of childhood abuse (scores above cut-off for at least two of the abuses).

Trauma events inventory. The trauma events inventory (TEI) is the primary measure of non-child abuse trauma in this study (41). Total numbers of different types of trauma among the participants was summed up for each category into a continuous variable. Although the TEI instrument also includes exposure to non-child abuse traumatic events in childhood, the mean age of exposure was 23.81 (0.70) y; hence, in the current study this refers to mostly adult trauma.

Sample Selection. Of the sample of 396 described above, 169 individuals who had experienced at least two types of moderate-to-severe childhood abuses or at least two other types of traumas during their lifetime were included in the analysis for this study. Of these, individuals with PSS $\leq 10$ and moderate-to-severe child abuse or two other types of traumas were selected as trauma-exposed individuals with neither current nor a history of lifetime PTSD symptoms (trauma controls) ($n = 108$). From a total of 61 individuals with current PTSD symptoms (PSS $> 20$ and significant symptoms in all PTSD B, C, and D clusters), we selected individuals (i) who had experienced at least two types of moderate-to-severe childhood abuses ($n = 32$) and (ii) who had experienced at least two other types of traumas but no child abuse ($n = 29$).
Statistical Analysis. Gene-expression data. Raw microarray scan files from Illumina HT-12 v3.0 arrays (Illumina) were exported using the Illumina Beadstudio program and passed back for further downstream analysis (www.R-project.org). Evaluation of the different microarray steps was done using the Illumina internal controls. Samples which were >5% SD were excluded. The data were transformed and normalized using the variance stabilizing normalization (42). A total of 15,877 probes passing the filter criteria of Illumina probe detection P value <0.01 in 5% of the individuals were used for subsequent analysis. To correct for confounding as a result of batch effects, the data were normalized using an empirical Bayes method for batch correction (43). Reproducibility of the gene-expression data were assessed using six pairs of technical replicates, yielding average Pearson correlations of 0.996. General linear models were constructed by regressing the gene-expression profiles against the PTSD group status and adjusting for sex, age, ethnicity, substance abuse, and treatment. The significance of association was estimated by two-tailed tests using R software (44). F test was calculated for multiple testing by 10,000 permutations using the permutation of regressor residuals test (http://cran.r-project.org/web/packages/glimperm/index.html). Briefly, the general linear models for each transcript were written as described above and the residuals of the regressions were permuted 10,000 times for each transcript using the shuffle-Z method to obtain the empirical P values corrected for multiple testing as described previously (34).

DNA-methylation data. Raw methylation Beta values from the HumanMethylation 450k BeadChip (Illumina) were determined via the Illumina Beadstudio program and loaded into R. Internal Illumina controls were used to assess the quality of staining, extension, hybridization, bisulfite conversion, and specificity. Samples with probe detection call rates <90% and those with an average intensity value of either <50% of the experiment-wide sample mean or <2,000 arbitrary units (AU) were excluded from further analysis, allowing 163 samples for subsequent analysis. Unsupervised hierarchical clustering was performed to identify extreme outliers and global trends in methylation. One sample of male DNA was included on each BeadChip as a technical control throughout the experiment and assessed for reproducibility, with average Pearson correlation coefficient of 0.993 across all replicates. Signals from methylated (M) and unmethylated (U) bead types were used to calculate a beta value as \( \beta = M/(U+M) \). Hybridization and chip batch effects were accounted for using an empirical Bayes method (43). The samples were quantile normalized and peak-corrected using the IMA package functions in R (44, 45). Hybridization and chip batch effects were accounted for using an empirical Bayes method (43). Methylation differences were calculated using generalized linear models in R by regressing the \( \beta \) values against the PTSD group and adjusting for sex, ethnicity, and substance abuse. Results were corrected for multiple testing by 10,000 permutations using the permutation of regressor residuals test (http://cran.r-project.org/web/packages/glimperm/index.html).

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the Childhood Trauma Questionnaire. Arch Gen Psychiatry 60(1):74–81.
ternal controls. Samples which were >5% SD were excluded. The data were transformed and normalized using the variance stabilizing normalization (42). A total of 15,877 probes passing the filter criteria of Illumina probe detection P value <0.01 in 5% of the individuals were used for subsequent analysis. To correct for confounding as a result of batch effects, the data were normalized using an empirical Bayes method for batch correction (43). Reproducibility of the gene-expression data were assessed using six pairs of technical replicates, yielding average Pearson correlations of 0.996. General linear models were constructed by regressing the gene-expression profiles against the PTSD group status and adjusting for sex, age, ethnicity, substance abuse, and treatment. The significance of association was estimated by two-tailed tests using R software (44). F test was calculated for multiple testing by 10,000 permutations using the permutation of regressor residuals test (http://cran.r-project.org/web/packages/glimperm/index.html). Briefly, the general linear models for each transcript were written as described above and the residuals of the regressions were permuted 10,000 times for each transcript using the shuffle-Z method to obtain the empirical P values corrected for multiple testing as described previously (34).

DNA-methylation data. Raw methylation Beta values from the HumanMethylation 450k BeadChip (Illumina) were determined via the Illumina Beadstudio program and loaded into R. Internal Illumina controls were used to assess the quality of staining, extension, hybridization, bisulfite conversion, and specificity. Samples with probe detection call rates <90% and those with an average intensity value of either <50% of the experiment-wide sample mean or <2,000 arbitrary units (AU) were excluded from further analysis, allowing 163 samples for subsequent analysis. Unsupervised hierarchical clustering was performed to identify extreme outliers and global trends in methylation. One sample of male DNA was included on each BeadChip as a technical control throughout the experiment and assessed for reproducibility, with average Pearson correlation coefficient of 0.993 across all replicates. Signals from methylated (M) and unmethylated (U) bead types were used to calculate a beta value as \( \beta = M/(U+M) \). Hybridization and chip batch effects were accounted for using an empirical Bayes method (43). The samples were quantile normalized and peak-corrected using the IMA package functions in R (44, 45). Hybridization and chip batch effects were accounted for using an empirical Bayes method (43). Methylation differences were calculated using generalized linear models in R by regressing the \( \beta \) values against the PTSD group and adjusting for sex, ethnicity, and substance abuse. Results were corrected for multiple testing by 10,000 permutations using the permutation of regressor residuals test (http://cran.r-project.org/web/packages/glimperm/index.html).