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Predictors of Inflammation in a Cohort of Bolivian Infants and Toddlers

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Abstract. Inflammation has been associated with cardiovascular disease and other health outcomes in children and adults, yet few longitudinal data are available on prevalence and predictors of inflammation in infants. We aimed to identify the prevalence of inflammation in a cohort of Bolivian infants and estimate its association with acute (recent illnesses) and chronic (overweight, stunting) morbidities and potential pathogen exposure (represented by water, sanitation, and hygiene [WASH] resources). We measured plasma concentrations of two acute phase proteins (C-reactive protein [CRP], marking acute inflammation, and alpha(1)-acid-glycoprotein [AGP], marking chronic inflammation) at three time points (target 2, 6–8, and 12–18 months). Of 451 singleton infants enrolled in the parent study, 272 had the first blood draw and complete data. Anthropometry and sociodemographic and recent illness data (2-week recall of cough, diarrhea, and fever) were collected at each visit. Inflammation was defined as CRP > 5 mg/L or AGP > 1 g/L.

The prevalence of inflammation increased from early infancy (3% at first blood draw) to later infancy (15–22% at later blood draws). Recent cough, recent fever, and age in months were significantly associated with relative increases in CRP (7–44%) and AGP (5–23%), whereas recent diarrhea was only significantly associated with an increase in CRP (48%). Neither anthropometry nor WASH was significantly associated with inflammation. Results confirm the role of recent acute illness in inflammation in infants, and indicate that adiposity and WASH are not as important to inflammation in this age category.

INTRODUCTION

Inflammation has been associated with numerous adverse health outcomes ranging from cardiovascular disease1–3 to psychiatric and mood disorders,4,5 to linear growth failure,6 of inflammation, to better understand the role of acute as

well as chronic exposures in this process, and identify points of intervention so as to prevent harmful consequences of inflammation. Adiposity is one known correlate of inflammation, and it is thought that adipose tissue is in fact pro-inflammatory; this association has been demonstrated in both children and adults.13–15 Socioeconomic status has also been frequently studied as a potential correlate of inflammation, with measures of increasing household wealth typically associated with reduced likelihood of elevated CRP.16–19

In developing-country settings, models of inflammation also often incorporate potential sources of pathogen exposure (and thus, infection), often represented by access to water, sanitation, and hygiene (WASH) resources such as where the family obtains water (and whether it is treated), the type of sanitation facilities, and the general cleanliness level of the house and surrounding area, with mixed findings.16,18–20

Indoor and outdoor pollutants have also been associated with inflammation in children and adults.20–22

Much research on correlates of inflammation has taken place in developed-country settings, which may not only have a different distribution of inflammation and risk factors, but also potentially a different relationship of risk factors to inflammation.19 Furthermore, some studies have suggested that prevalence and correlates of inflammation may differ by sex as well as by age.16,18,19,23,24

Although a number of studies have examined inflammation in schoolchildren, and others have described extreme states such as sepsis in neonates, very few studies have been published on prevalence or correlates of inflammation in healthy infants in community settings, particularly in developing countries where inflammation may be prevalent.16,25 Nonetheless, this is an important population to study given the potential adverse consequences of inflammation and immune dysregulation even in infants.26,27

The present study leverages data from a longitudinal birth cohort of Bolivian infants to identify the prevalence and

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correlates of inflammation in young infants and examine how these may change with increasing infant age. Bolivia, as a lower-middle-income country in the midst of the epidemiologic transition and suffering the “double burden” of malnutrition and overweight, is a location well suited to studying varied contributors to inflammation. This work will focus on the role of recent illness, anthropometry, and pathogenic exposures (represented by access to WASH resources); these factors have been shown to be correlated with inflammation in adults and schoolchildren, but there are very few data in infants. The present work will help to elucidate different inflammatory factors in a community population of healthy infants in a developing-country urban setting. Furthermore, it will add to knowledge regarding correlates of inflammation in infants, whose innate immune function differs from that of adults. Also, nearly all studies of correlates of inflammation have only taken into account CRP, a marker of the earlier stages of inflammation. In the present study, we also add AGP, a marker of later stages of inflammation, as a way to better understand how correlates may differ by acute versus chronic inflammatory processes.

**METHODS**

**Study population and design.** Data for the present study were drawn from the *Nutrición, Inmunología, y Diarrea Infantil* study, the primary aim of which was to assess the effect of nutritional status on response to the rotavirus vaccine. For this reason, all enrolled infants were required to receive both doses of the rotavirus vaccine (scheduled at 2 and 4 months of age). In brief, 461 healthy infants (2–4 weeks of age) and their mothers were recruited from two hospitals in El Alto, Bolivia (altitude 4,000 m), during well-child or vaccination visits. This is a primarily urban and largely indigenous population; although most have access to sanitation and improved water, socioeconomic resources are typically low. Exclusion criteria included infant illness at recruitment, infants suspected to have immunodeficiency (e.g., human immunodeficiency virus), congenital malformations, and maternal inability to speak and understand Spanish or Aymara. Recruitment took place from May 2013 to March 2014, and infant–mother dyads were followed for 12–18 months, with final data collected in March 2015. The study comprised seven hospital visits and two in-home visits, with blood drawn from mothers at a target schedule of 1 and 6–8 months postpartum, and from infants at a target schedule of 2, 6–8, and (optionally) 12–18 months of age. However, despite interviewer efforts, some blood draws fell outside the target ranges due to participant family travel or faulty contact information. The third blood draw, added to support a newly funded secondary aim, was approved only after 50% of infants had already completed the study; these mothers were contacted to participate optionally in an extra visit and blood draw. Although infants were required to be healthy at the first visit (0–2 months of age), no infants were excluded from blood draws based on concurrent or recent illness. For some infants, the first blood draw coincided with Rotarix vaccine, but no association with inflammatory status was detected. For the present study, only singleton infants who had completed at least the first blood draw (*N* = 365) were eligible, and of these, 272 had nonmissing data for exposures and outcomes (Figure 1).

**Ethical approval.** The protocol and instruments for this study were approved by the Emory University Institutional Review Board (IRB00056127) and the Bolivian “Comité de Etica de la Investigación” (Research Ethics Committee). Mothers provided written informed consent in Spanish or Aymara.

**Laboratory analysis and definitions of inflammation.** Venous blood was collected by trained phlebotomists using sterile, disposable equipment. Blood was drawn into zinc-free microtainers using butterfly needles, and stored at 2–8°C before transport later that day to our partner laboratory in La Paz, Bolivia. Samples were then frozen before shipping to Emory, where they were aliquoted and shipped for further analyses. Plasma was analyzed by sandwich enzyme-linked immunosorbent assay for CRP (a marker of inflammation; limit of detection [LOD] = 0.5 mg/L) and AGP (a marker of inflammation; LOD = 0.1 g/L). Quality control values for CRP and AGP indicated high quality at lower (CRP < 0.1 mg/L; AGP < 0.1 g/L) and higher (CRP up to 45 mg/L, AGP up to 2.5 g/L) ranges.

**Data collection.** Sociodemographic data were collected by trained Bolivian interviewers at the first study visit via questionnaire. At each visit, interviewers collected data on maternal report of infant morbidities and feeding practices over the 2 weeks before the interview. Anthropometry was conducted by a two-person team of trained interviewers. Infants were weighed nude or with light clothing (no diaper) on a Seca scale (weight measured to the nearest 0.1 kg), and measured on a ShorrBoard (length measured to the nearest 0.1 cm). Weight-for-length and length-for-age Z-scores (WLZ and LAZ, respectively) were calculated based on World Health Organization (WHO) references and using the WHO SAS macro. In the field, stunted (LAZ < −2) and wasted (WLZ < −2) infants were identified by study staff according to WHO growth charts and were referred as appropriate.

**Variable definitions for outcomes, exposures, and covariates.** Inflammation was defined as CRP > 5 mg/L or AGP > 1 g/L.9,32 As described above, acute illness, anthropometry, and WASH resources were considered as primary exposures of interest. Acute illnesses were defined as positive 2-week maternal recall of infant symptoms; diarrhea, fever, and respiratory symptoms were included in each model as indicator variables. Anthropometric variables included overweight or obesity (WLZ > 1) and stunting (LAZ < −2). Several measures of WASH resources (a proxy of pathogenic exposure) were considered: sewer type (piped versus other, e.g., pit latrine), private toilet (versus shared or no access), water source (piped indoors versus other, e.g., piped outdoors), water treatment (any versus none), trash disposal (picked up versus other, e.g., thrown onto patio), crowding (> 2 people sleeping per bedroom). Potential covariates included preterm birth (< 37 completed gestational weeks), caesarian section (versus vaginal) birth, months of exclusive breastfeeding completed at the time of visit (defined as months during which infant had received only breast milk, with no other liquids or solids), maternal employment, maternal education (categorized as primary or less, at least some secondary, or at least some superior [reference]), and sample-specific wealth index (created via principal components analysis of assets and house materials, divided into quintiles with the highest quintile as reference23). Cow’s milk intake was
considered for the models, but could not be included due to low prevalence (0% at 2 months, though 14–30% at later visits). Concurrent or recent vaccination was not considered as a confounder, given its potential role as an intermediate between age and inflammation.

**Statistical methods.** Given that the same infants were followed over time and contributed multiple measurements, methods appropriate for clustered data were applied. AGP and CRP were considered first separately and as continuous variables, and then as categorical variables (elevated/nonelevated), and finally together (dichotomized inflammation defined as either elevated AGP or CRP). Potential confounders of the relationships of acute illnesses, sanitation, and obesity, to AGP and CRP were selected a priori based on literature review and directed acyclic graph analysis, and were then included in the initial models based on significant bivariate association with the exposure or the outcome.

Continuous AGP and CRP were log transformed (base 10) to meet normality assumptions, and linear models were used to test relationships of predictors to the outcome; percent relative changes in the outcome are reported for each predictor. Percent relative change was calculated as $100 \times \left(10^\beta - 1\right)$, where $\beta$ is the coefficient associated with predictor X and thus $10^\beta$ is equal to the ratio of $E(Y \mid X = 1)$ to $E(Y \mid X = 0)$. For categorized AGP and CRP, logistic regression was used. In each case, mixed models were first applied, and a random intercept was included if significant (approximate likelihood ratio test). If nonsignificant, generalized estimating equations (GEE) were fit using an appropriate distribution (binomial for dichotomized outcomes) and an exchangeable correlation structure. Interactions of acute illnesses with time were tested using likelihood ratio tests (with maximum likelihood estimates) for mixed models and Wald tests for GEE models, and $P < 0.05$ was considered significant. Collinearity was assessed for each model using condition indices and variance decomposition proportions, and models reduced as necessary until collinearity was no longer present. Final linear mixed models used restricted maximum likelihood estimation and were fit using the lme4
package in the R Environment for Statistical Computing (R Core Team, Vienna, Austria). Marginal $R^2$ (reflecting the proportion of variance explained by fixed factors) and conditional $R^2$ (reflecting the proportion of variance explained by fixed and random factors) were calculated for linear mixed models using the piecewise SEM R package (http://arxiv.org/abs/1509.01845). Fixed effects were tested using F tests with the Satterthwaite approximation for denominator degrees of freedom. GEE models (http://CRAN.R-project.org/package=gee) were fit using the R gee package, and Wald tests based on robust variances are presented for fixed effects. Data were cleaned and analyzed using SAS v9.4 (Cary, NC) and the R Environment for Statistical Computing.

RESULTS

Characteristics of the study sample. Of 451 enrolled singletons, 365 infants (80.9%) had at least the first blood draw (target schedule 2 months), and 272 of these (74.5%) had nonmissing data for exposures and outcomes. The most common missing data were for wealth index (16% missing) and type of sewer (8% missing). Of these 272, 239 (87.8%) also completed a second blood draw (at a target schedule of 6–8 months), whereas 126 infants (46.3%) completed a blood draw at the final study visit (target schedule of 12–18 months), and 121 had all three blood draws. Although some blood draws took place outside the target age range, the majority were within the target ranges (Figure 2).

Infants were fairly evenly distributed in terms of gender, nearly one-third were born via caesarian section, and one-fifth were born preterm (Table 1). At least 60% of the sample had access to improved WASH for at least one measure, though only 40% had water that was piped inside their home (the majority of the rest had access to water that was piped to a point outside their home but close by, often in the same lot). Mothers had a mean age of 26 years, one-quarter were employed, and most had at least some secondary education. About one-quarter of households owned a refrigerator, and one-third had high-quality flooring material. The prevalence of exclusive breastfeeding was 60% at 1–5 months and declined with age (Table 2). Stunting varied between 14% and 18%, whereas overweight varied from 32% at the first blood draw to 19% at the third blood draw.

Presence of inflammation and recent illness. Both CRP and AGP were right skewed at all time points, with both also right shifted at the second and third blood draws (at approximately 6–10 and 10–18 months of age) as compared with the first blood draw (at approximately 2 months of age; Figure 3). Almost no inflammation (elevated CRP or AGP) was detected at the first blood draw when infants were 1–5 months of age (3%), but the prevalence of inflammation rose to 22% among infants assessed at 6–10 months of age and 15% among the infants assessed at 10–18 months.

The prevalence of recent illness among infants varied by age group, with younger infants tending to have lower prevalence of recent illness (Table 2). Specifically, the prevalence of diarrhea increased with increasing age from 14% among infants 1–5 months to 24% among infants 10–18 months of age. Cough or respiratory illness was highly prevalent, with 40% of the youngest infants affected and around 50% of the

![Figure 2](image-url)
older age groups affected. The prevalence of fever was also high, with 21% among the youngest infants and around 40% for older infants and toddlers.

**Results of regression models.** Alpha(1)-acid-glycoprotein. Continuous (log-transformed) infant AGP was significantly associated with recent cough (relative increase of 18%, \( P = 0.0001 \)), recent fever (relative increase of 23%, \( P < 0.0001 \)), and age in months (relative increase of 5% per month, \( P < 0.0001 \); Table 3). Marginal and conditional \( R^2 \), representing the proportion of variance explained by fixed, and fixed plus random effects, respectively, were 27% (marginal) and 35% (conditional). Elevated AGP (> 1 g/L) was also significantly associated with recent cough (odds ratio [OR] = 3.9, \( P = 0.001 \)), recent fever (OR = 2.8, \( P = 0.013 \)), and age in months (OR = 1.1, \( P < 0.0001 \), for a 1-month increase; Table 4). Interactions between acute illnesses and age were tested for each model and found to be nonsignificant (data not shown).

C-reactive protein. Continuous (log-transformed) infant CRP was significantly associated with recent diarrhea (relative increase of 48%, \( P = 0.01 \)), recent cough (relative increase of 44%, \( P = 0.003 \)), recent fever (relative increase of 35%, \( P = 0.021 \)), and age in months (relative increase of 7% per month, \( P < 0.0001 \); Table 3). Marginal and conditional \( R^2 \) were 13% and 29%, respectively. Elevated CRP (> 5 mg/L) was significantly associated with recent diarrhea (OR = 2.2, \( P = 0.015 \)) and age in months (OR = 1.1, \( P = 0.001 \), for a 1-month increase), but not with other recent illnesses (Table 4). Interactions between acute illnesses and age were tested for both models and found to be nonsignificant (data not shown).

Inflammation. Dichotomized infant inflammation (elevated AGP or CRP) was significantly associated with recent fever (OR = 2.1, \( P = 0.019 \)) and age in months (OR = 1.1, \( P < 0.0001 \), for a 1-month increase; Table 4). Interactions between acute illnesses and age were tested and found to be nonsignificant (data not shown).

**DISCUSSION**

In this prospective cohort study of Bolivian infants, the prevalence of any inflammation increased from 3% at the first assessment (approximately 2 months of age) to 15–22% at later time points (approximately 6–8 and 12–18 months of age for most infants). Although this is much lower than the prevalence of elevated CRP (> 5 mg/L) in a cohort of Tanzanian infants 6–35 months of age,16 the overall geometric mean CRP in our older infants (0.80 mg/L at the third blood draw, ~12–18 months of age) is comparable to, though somewhat higher than, that in a cohort of European and south Asian British infants at 12 months of age (0.69 mg/L for European-origin infants, 0.51 mg/L for south Asian-origin infants).25 These comparisons potentially reflect population differences in socioeconomic status, rurality, age, and other correlates or

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**TABLE 1**

Characteristics of the study sample \((N = 272)\)

<table>
<thead>
<tr>
<th>Infant characteristics</th>
<th>Frequency* or mean (±SD) %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (days) at enrollment</td>
<td>34 ± 8</td>
</tr>
<tr>
<td>Male</td>
<td>144 52.9</td>
</tr>
<tr>
<td>C-section</td>
<td>81 29.8</td>
</tr>
<tr>
<td>Preterm (&lt; 37 weeks gestational age)</td>
<td>48 17.6</td>
</tr>
<tr>
<td>Water piped indoors</td>
<td>108 39.7</td>
</tr>
<tr>
<td>Treats water before drinking</td>
<td>176 64.7</td>
</tr>
<tr>
<td>Private toilet (vs. shared or none)</td>
<td>170 62.5</td>
</tr>
<tr>
<td>Piped sewer (vs. pit latrine)</td>
<td>226 83.1</td>
</tr>
<tr>
<td>Trash picked up (vs. thrown onto patio, other)</td>
<td>226 83.1</td>
</tr>
<tr>
<td>Crowding (&gt; 2 people per bedroom)</td>
<td>200 73.5</td>
</tr>
<tr>
<td>Water and sanitation</td>
<td></td>
</tr>
<tr>
<td>Maternal age (years)</td>
<td>25.5 ± 6.3</td>
</tr>
<tr>
<td>Maternal employment</td>
<td>70 25.8</td>
</tr>
<tr>
<td>Household owns refrigerator</td>
<td>77 28.3</td>
</tr>
<tr>
<td>Higher quality floor material (hardwood, carpet, or tile vs. cement)</td>
<td>98 36.0</td>
</tr>
<tr>
<td>Maternal education</td>
<td></td>
</tr>
<tr>
<td>Primary or less</td>
<td>35 12.9</td>
</tr>
<tr>
<td>At least some secondary</td>
<td>167 61.4</td>
</tr>
<tr>
<td>At least some superior</td>
<td>70 25.7</td>
</tr>
</tbody>
</table>

\*Of singleton infants with plasma available for at least the 2-month visit and no missing data for exposures or outcomes.
**Table 3**
Mixed linear regression models of associations of predictors and (log-transformed) AGP and CRP

<table>
<thead>
<tr>
<th>AGP* (N = 626 observations of 272 infants)</th>
<th>CRP† (N = 626 observations of 272 infants)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percent change‡</td>
<td>95% CI</td>
</tr>
<tr>
<td>Recent illness (2-week recall) and age</td>
<td></td>
</tr>
<tr>
<td>Diarrhea</td>
<td>7.6</td>
</tr>
<tr>
<td>Cough or respiratory illness</td>
<td>17.7</td>
</tr>
<tr>
<td>Fever</td>
<td>23.3</td>
</tr>
<tr>
<td>Age (1-month increase)</td>
<td>4.9</td>
</tr>
<tr>
<td>Anthropometry</td>
<td></td>
</tr>
<tr>
<td>Stunted (LAZ &lt; -2)</td>
<td>-2.0</td>
</tr>
<tr>
<td>Overweight (WLZ &gt; 2)</td>
<td>1.3</td>
</tr>
<tr>
<td>Water and sanitation</td>
<td></td>
</tr>
<tr>
<td>Water piped indoors (vs. other source)</td>
<td>7.1</td>
</tr>
<tr>
<td>Treats water</td>
<td>5.5</td>
</tr>
<tr>
<td>Private toilet (vs. shared with other households or none)</td>
<td>6.1</td>
</tr>
<tr>
<td>Flush toilet (vs. other)</td>
<td>7.5</td>
</tr>
<tr>
<td>Trash picked up (vs. burned or otherwise disposed of)</td>
<td>-4.6</td>
</tr>
<tr>
<td>Crowding (&gt; 2 people per bedroom)</td>
<td>-2.5</td>
</tr>
<tr>
<td>Other sociodemographics</td>
<td></td>
</tr>
<tr>
<td>Preterm birth (&lt; 37 weeks gestational age)</td>
<td>7.8</td>
</tr>
<tr>
<td>Months of exclusive breastfeeding (1-month increase)‡</td>
<td>-0.6</td>
</tr>
<tr>
<td>Maternal education</td>
<td></td>
</tr>
<tr>
<td>Primary or less</td>
<td>0.7</td>
</tr>
<tr>
<td>At least some secondary</td>
<td>1.5</td>
</tr>
<tr>
<td>At least some superior (reference)</td>
<td>0.0</td>
</tr>
<tr>
<td>Wealth index**</td>
<td></td>
</tr>
<tr>
<td>First (lowest) quintile</td>
<td>9.8</td>
</tr>
<tr>
<td>Second quintile</td>
<td>6.3</td>
</tr>
<tr>
<td>Third quintile</td>
<td>10.0</td>
</tr>
<tr>
<td>Fourth quintile</td>
<td>4.5</td>
</tr>
<tr>
<td>Fifth (highest) quintile (reference)</td>
<td>0.0</td>
</tr>
</tbody>
</table>

AGP = alpha1(-)-acid-glycoprotein; CRP = C-reactive protein; LAZ = length-for-age Z-score; SD = standard deviation; WLZ = weight-for-length Z-score.

* All available observations with nonmissing data included for all infants with at least first blood draw. Random intercept with variance 0.004 was significant with P = 0.013 (approximate likelihood ratio test). The marginal R^2, reflecting variance explained by fixed effects, was 0.27, and the conditional R^2, reflecting variance explained by fixed and random effects, was 0.35.

† All available observations with nonmissing data included for all infants with at least first blood draw. Random intercept with variance 0.065 was significant with P = 0.0004 (approximate likelihood ratio test). The marginal R^2, reflecting variance explained by fixed effects, was 0.13, and the conditional R^2, reflecting variance explained by fixed and random factors, was 0.29.

‡ Percent relative change calculated as 100 × (10^p – 1), where p is the coefficient associated with predictor X and thus 10^p is equal to the ratio of E(Y | X = 1) to E(Y | X = 0).

§ F tests with Satterthwaite approximation for degrees of freedom.

¶ Months of exclusive breastfeeding defined as total months, at time of assessment, that infant was fed only with breast milk, with no other liquids or food. One month of formula was allowed as long as breastfeeding was concurrent.

** Constructed using principal component analysis of household assets and construction materials.
potential modifiers of inflammation. Our study also measured AGP and found it to be elevated in only 2% of infants at the first blood draw, but > 10% of infants at later blood draws. This is a new finding, as previous studies of inflammation in healthy infants have not measured AGP.

In our study of Bolivian infants, inflammation was significantly associated with maternal recall of recent illnesses and increased age in months regardless of biomarker. The association of recent illness with inflammation and the larger magnitude of the effect in linear CRP models as opposed to linear AGP models are consistent with the biology of the APR. Specifically, CRP experiences a much more dramatic relative increase as compared with AGP during the APR.6

We did not see significant interactions of acute illness and age, which may suggest that the infant inflammatory reaction to infectious illness does not change significantly between 2 and 18 months of age. Although two other identified studies demonstrated a significant effect of age in the opposite (negative) direction, these studies included older children: one compared children at 6–35 months to those at 36–59 months,16 whereas the other included only children 2–15 years of age.18 It is possible that age is not linearly related to inflammation. Possibly, inflammation in developing-country settings peaks between 12 and 24 months of age, and then declines as children’s immune systems are trained. Another study in British infants more similar in age to our study population did not show any significant effect of age, but took place in a low-inflammation, high-resource, developed-country setting.25 We hypothesize that our finding of increasing inflammation with age may be related to the fact that older infants and toddlers will have much more opportunity for inflammatory exposures (e.g., pathogens) given their increased mobility as well as the incorporation of new foods and potentially unsafe liquids during weaning. Although it is possible that the prevalence of inflammation could be affected by the receipt of vaccines concurrent with blood draws, we did not see a significant relationship between receipt of Rotarix® and inflammation at the first blood draw (data not shown). Furthermore, while older infants (~12 months) were more likely than younger infants (~6 months) to receive more vaccines near the time of the blood draw (due to the vaccine schedule), the prevalence of inflammation was similar between these two groups. Therefore, we hypothesize that our findings of increased inflammation at older ages are unlikely to be explained by variations in vaccination. Although our models did not demonstrate a significant effect of cumulative months of exclusive breastfeeding, a sensitivity analysis assessing current breastfeeding practices within each age group did show a trend toward a protective effect of exclusive breastfeeding among younger infants (data not shown). A study in Tanzanian children also suggested a “protective” effect of breastfeeding.16
Neither marker of anthropometry (stunting nor overweight, both at low-to-moderate levels in the population) was significantly associated with inflammation in this cohort. Although an association of adiposity and inflammation has been well established in adults, and other authors have found similar associations in children, we did not see a significant association of WLZ with inflammation in our cohort. Alternative definitions of adiposity (using body mass index–for–age Z–score as well as cutoffs for obesity versus normal–overweight) also failed to yield significant results in our models, suggesting that differences in adiposity definition were not the explanation for our results. Our results may be due to a lack of power or to the lower prevalence of overweight in our cohort as compared with developed–country children. Alternatively, it may be that our infants were too young for the inflammatory effect of adiposity to have presented; the studies showing significant associations were all in children of at least preschool age. Although a study of Ecuadorian children found a negative association between CRP and attained growth in children 2–7 years of age, two other studies of inflammation that included infants (one in British infants 3–24 months of age, one in Tanzanian children 6–59 months of age) did not find significant associations between body size measures, including length, and CRP.

Markers of WASH resources were also not significantly associated with individual markers of inflammation in bivariate or multivariable analyses. Though this was contrary to expectations, as we hypothesized that lack of access to WASH resources would reflect increased pathogen exposure, it is consistent with results from Hadley and others and Thompson and others, who also failed to see significant associations of WASH resources (e.g., private toilet, water quality) in multivariable models of child inflammation. It may be that the available measures of WASH were not granular or specific enough to capture true pathogen exposure, or that young infants are more protected from these effects given their lower mobility, restricted eating, and close maternal supervision. It may also be that WASH measures are more important to a longer–term process of immune development as opposed to short–term markers of inflammation in young infants, whose immune systems may still be undeveloped.

Calculated $R^2$ for linear models indicated that fixed effects explained a larger proportion of AGP variance as compared with CRP variance, but that no more than a third of variance was explained by the parameters in either model. This suggests that factors outside of our measured morbidities, anthropometry, and WASH resources may be important to the development of inflammation in infants. The similar $R^2$ (marginal and conditional) for the model of AGP implies that within–subject correlations are less important (supported by the small magnitude of the variance of the random intercept term), and that other environmental factors or behavioral factors may be important. Conversely, the large jump from the marginal to the conditional $R^2$ for the model of CRP suggests that within–subject factors are important for CRP levels, potentially implying constitutional differences or other factors present from birth. To our knowledge, this is the first that this type of analysis has been applied to the study of inflammation in healthy infants.

Although the inflammation cutoff of 5 mg/L for CRP is widely used, it has been suggested that this cutoff is inappropriate for healthy or pediatric populations. Therefore, we reran the model of elevated CRP with a cutoff of > 3 mg/L, which has previously been suggested and used as a “high–risk” cutoff. This analysis demonstrated similar results in terms of significance to the model with the original cutoff (5 mg/L), although the effect of recent diarrhea became nonsignificant, whereas recent cough became significant (OR = 2.2 [1.2–4.0], $P = 0.012$). We also tested a cutoff of > 1.1 mg/L as suggested by Wander and others’ analysis of Tanzanian children. This model also showed similar results to the models with cutoffs of 3 and 5 mg/L. However, recent fever (OR = 1.8 [1.2–2.7], $P = 0.005$) and trash picked up (OR = 0.6 [0.3–1.0], $P = 0.043$) became significant (recent diarrhea was again nonsignificant). We could not perform an analysis using a cutoff of CRP > 10 mg/L (often suggested as a marker of acute inflammation or infection), because not enough infants at all time points exhibited CRP values this high.

This study has several strengths. First, we were able to follow infants over the course of their 1st year of life, gathering data on inflammation and its potential correlates at multiple time points. This enabled us to test whether the effects of acute illness varied over time as infants’ immune systems matured. We also had access to a rich variety of data, including not only recent illness recall, but also anthropometry, birth characteristics, WASH resources, and other sociodemographics. Furthermore, we were able to test two separate markers of inflammation—CRP and AGP—that reflect different time points in the APR. However, our study is limited by our decreased sample size at the third visit, which negatively impacts our power to test associations with many predictors at once. Comparison of the full analytical sample with the subset with data at all three visits revealed similar prevalence of elevated inflammatory biomarkers (though cough and diarrhea were slightly less prevalent among younger infants) and few differences in characteristics—the prevalence of preterm infants was slightly lower in the subset (13% versus 18%), whereas the prevalence of overweight among 2–month–olds was slightly lower (24% versus 32%) and mothers were slightly more educated (32% with superior education versus 26%). Linear and logistic models for both AGP and CRP gave similar conclusions with the subset as compared with the analytic dataset, though the significance of some results varied (the effects of flushing toilet and preterm birth became significant, but with a wide confidence interval, in several models, whereas the effects of wealth index and water treatment became significant in the logistic model of CRP, and the effect of fever became nonsignificant in two models; data not shown). Although this may point to a mild selection bias resulting from dropout at the third visit, it is reassuring that our reported results are in the same direction and toward the null as compared with those in the subset. Another possible limitation was that the reported LOD for the CRP assay used was higher than that of some high–sensitivity CRP assays, and may not fully capture low levels of inflammation. However, this higher LOD does not affect the main conclusions of this study. Another limitation is that we did not have information on other potential markers of pathogen exposure, such as the presence of feces around the home, and that morbidities were...
captured via maternal recall and not verified by examination in all cases.

CONCLUSIONS

In this study, we found that inflammation increased from early infancy to later infancy and early toddlerhood, potentially reflecting increased exposure to pathogens and other inflammatory agents in this cohort of Bolivian infants. We also found that recent illnesses were significantly associated with CRP and AGP, but did not fully explain differences in either acute phase protein. This cohort of infants did not demonstrate significant associations between inflammation and WASH or anthropometry, suggesting that these exposures may be more relevant to older children. Overall, the results underscore the importance of biochemical measures of inflammation in infants, given that inflammation cannot be identified by sociodemographic or morbidity information alone, and may manifest only subclinically.

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