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OY Addo, Emory University
Aryeh D Stein, Emory University
CHD Fall, University of Southampton
DP Gigante, Universidade Federal de Pelotas
AM Guntupalli, University of Southampton
BL Horta, Universidade Federal de Pelotas
CW Kuzawa, Northwestern University
N Lee, University of San Carlos
SA Norris, University of the Witwatersrand
C Osmond, University of Southampton

Only first 10 authors above; see publication for full author list.

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Parental Childhood Growth and Offspring Birthweight: Pooled Analyses from Four Birth Cohorts in Low and Middle Income Countries

O.Y. ADDO,1,* A.D. STEIN,1 C.H.D FALL,2 D.P. GIGANTE,3 A.M. GUNTUPALLI,4 B.L. HORTA,3 C.W. KUZAWA,5 N. LEE,6 S.A. NORRIS,7 C. OSMOND,8 P. PRABHAKARAN,9 L.M. RICHTER,1 H.P.S. SACHDEV,7 AND R. MARTORELL,1
AND ON BEHALF OF THE COHORTS GROUP

Objective: Associations between parental and offspring size at birth are well established, but the relative importance of parental growth at different ages as predictors of offspring birthweight is less certain. Here we model parental birthweight and postnatal conditional growth in specific age periods as predictors of offspring birthweight.

Methods: We analyzed data from 3,392 adults participating in four prospective birth cohorts and 5,506 of their offspring.

Results: There was no significant heterogeneity by study site or offspring sex. 1SD increase in maternal birthweight was associated with offspring birthweight increases of 102 g, 1SD in maternal length growth 0–2 year with 46 g, and 1SD in maternal height growth Mid-childhood (MC)-adulthood with 27 g. Maternal relative weight measures were associated with 24 g offspring birth weight increases (2 year-MC) and 49 g for MC-adulthood period but not with earlier relative weight 0–2 year. For fathers, birthweight, and linear/length growth from 0–2 year were associated with increases of 57 and 56 g in offspring birthweight, respectively but not thereafter.

Conclusions: Maternal and paternal birthweight and growth from birth to 2 year each predict offspring birthweight. Maternal growth from MC-adulthood, relative weight from 2-MC and MC-adulthood also predict offspring birthweight. These findings suggest that shared genes and/or adequate nutrition during early life for both parents may confer benefits to the next generation, and highlight the importance of maternal height and weight prior to conception. The stronger matrilineal than patrilineal relationships with offspring birth weight are consistent with the hypothesis that improving the early growth conditions of young females can improve birth outcomes in the next generation. Am. J. Hum. Biol. 27:99–105, 2015.

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Poor nutrition in early life may influence fetal and infant growth of future generations via genetic, epigenetic, physiological, and other mechanisms. Improved nutrition prior to age 2 year has been linked to taller adult stature, improved cognitive and school performance and in males, only, increased economic productivity and earnings (Hod- dinott et al., 2008, 2013; Martorell et al., 2010; Stein et al., 2010); these characteristics can influence offspring birthweight by improving, among others, household welfare and the condition of women. Nutritional supplementation in Guatemalan girls was associated with substantial increases in their offspring’ birthweight and height and head circumference in childhood; supplementation in boys was not associated with offspring’ growth (Behrman et al., 2009). The size of the uterus and the ovaries are reduced in women born with intrauterine growth restriction, which can influence reproductive outcomes (Ibanez et al., 2003). Other mechanisms for intergenerational influences are also possible and are not mutually exclusive. These include shared environments and genes as well as epigenetic changes resulting from nutritional exposures during embryogenesis, intra-uterine and early postnatal life (Burdge et al., 2013; Martorell and Zongrone, 2012; Pem- brey et al., 2013).

Past studies have demonstrated associations of parental anthropometry birthweight (and adult height and Body Mass Index, BMI) with offspring birth size, with most but not all finding evidence for stronger matrilineal than patrilineal inheritance of birthweight, BW (Addo et al., 2013; Kuzawa and Eisenberg, 2012; Mattsson and...
Adult height and BMI are cumulative measures and little is known about the relative importance for offspring birthweight of parental growth at specific periods during childhood. Such information would be useful for identifying critical periods during which growth failure has greater consequences and for setting priorities for nutrition interventions. To our knowledge only one study, from Pelotas, Brazil, has related growth during childhood in mother and fathers to offspring birthweight (Horta et al., 2009). The authors reported that mothers who gained weight faster than the population mean from birth to 20 months, but not from 20 months to 42 months, had heavier firstborn babies. In contrast, paternal weight gain at these ages was unrelated to offspring BW.

Here we build upon past work in this area by including three cohorts in addition to the Pelotas cohort and by examining all births, not just first-borns, while also including measures of growth in height and weight from birth through adulthood. Specifically, we assess the relative importance of parental birthweight and growth during different growth intervals from birth to adulthood as predictors of offspring birthweight in four cohorts from low and middle income countries (LMICs).

Study methods

The data for this study come from 4 of the 5 birth cohorts that participate in the Consortium on Health Oriented Research in Transitional Societies (COHORTS) (Richter et al., 2012). We use data from the 1982 Pelotas Birth Cohort–Brazil (Victora and Barros, 2006), the Institute of Nutrition of Central America and Panama Nutrition Trial Cohort–Guatemala (Stein et al., 2008), the New Delhi Birth Cohort–India (Sachdev et al., 2005) and the Cebu Longitudinal Health and Nutrition Survey CLHNS–Philippines (Adair et al., 2011). The Birth to Twenty Cohort (BT20)–South Africa (Richter et al., 2007), also a member of COHORTS, is the youngest, with only a small numbers of births to date and was not included in this analysis. Attrition information by site for the last follow-up was provided elsewhere (Richter et al. 2012). The number of people not traced into adulthood varied between 1% in the Philippines and 17% in Brazil. Other causes of attrition were death, usually in early childhood, and migration out of the area or country. The percentage of the original cohort members examined at the last follow-up was over 65% in three cohorts (Brazil, Guatemala, and the Philippines) and 29% in India, where nearly half the original participants moved out of the study area and could not be examined at follow-up. Prior analyses of COHORTS data indicate that estimated coefficients do not seem to have been biased by attrition or missing data (Adair et al., 2013). Finally, between 41 and 83% of adults examined at last follow-up (with information on birthweight and height and weight at other periods including adulthood) provided birthweight information for at least one child.

Data collection

Parental measures. Parental birth weights and postnatal anthropometry were measured following site-specific protocols by trained observers. Birthweight was measured in hospitals at delivery in Brazil, in the community within 72 h of birth in India, in hospitals or at home by birth attendants in The Philippines and by a project nurse at home or in a health post in Guatemala. Mid childhood (MC) heights and weights were measured at variable ages due to site-specific differences in data collection schedules. It was at 4 years in Brazil, Guatemala and India, and 5.5 years in the Philippines. We used growth variables of each parent as study exposures. We defined maternal and paternal childhood stunting using the WHO height-for-age z scores (HAZ) ≤−2SD cut-off at 2 year (WHO Multicentre Growth Reference Study Group, 2006).

Offspring measures. Offspring birth weights were reported by the parent, except in Guatemala where they were obtained by measurement for births before 1999 and by recall later (Behrman et al., 2009; Ramakrishnan et al., 1999). Offspring birthweight and low birthweight (LBW, <2,500 g) were considered as study outcomes.

Other variables

We included several variables as covariates: parental early childhood household socioeconomic status (quintiles, lowest as reference category), parental birth order (categorized into first-borns and non-firstborns), maternal age at delivery of offspring (categorized as <18y or ≥18y), type of nutrition supplementation received in Guatemala, and offspring sex. Parental early childhood SES was represented by father’s income in Brazil, father’s occupation and mother’s education in India, and by asset scores, derived from principal component analyses (PCA) of household items for all other sites. As the parents in the Guatemala cohort had participated as children in a nutrition supplementation trial that improved linear growth (Habicht et al., 1995), we created a three category variable (one for each of the Guatemala 2 trial arms, and the 3rd category for the other sites that had no intervention). In pooled mixed models we also controlled for study site.

Data management

Analytic sample and inclusion criteria. The primary inclusion criterion for the analytic study sample was availability of birthweight for the parent and the child (n = 4,906). Hundred and seven (1.9%) pairs were excluded because the child was not a singleton.

Parental conditional growth. We used conditional height and conditional relative weight variables to assess growth (length/height and weight) from birth through adulthood (Adair et al., 2013; Keijzer-Veen et al., 2005; Li et al. 2003). The conditional variables are residuals from site- and sex-specific linear regression models. Conditional height is present length or height accounting for previous length or height, and weight (but not present weight). Conditional relative weight is present weight accounting for present height and all previous weight and height measures. For example, adult conditional relative weight is derived from a regression of adult weight on adult height, weight and height at mid-childhood, weight and length when aged 2 years, and birthweight. The conditions may be interpreted as the deviation from the child’s own predicted growth trajectory in relation to that of its peers, and hence measure relative growth within an
age interval (Adair et al., 2013). Conditional growth variables were expressed in SD units to allow for comparisons of coefficients across measures, periods and parental sex. As conditional height and relative weight variables are by definition uncorrelated, they can be included in a single multivariate model without causing variance inflation. The conditional variables were estimated for the periods: birth-2 year, 2 year-Mid-childhood (MC) and MC-adulthood. For convenience, we name these periods early, middle, and late childhood periods of maternal growth. Because birth length was not available for Brazil, conditional variable derivations used birthweight as a measure of birth size for all sites. Conditional height refers to growth in length/height while conditional relative weight refers to weight gain independent of linear growth.

**Statistical methods**

Descriptive characteristics are presented as means (SD) for continuous variables and as percentages for categorical variables. We estimated associations of parental birthweight and conditional heights and relative weights with offspring birthweight, separately for mothers and fathers, using mixed linear models adjusted for parental level covariates, offspring sex and study site as fixed effects. As we found significant intra family clustering of offspring birthweights (intra cluster correlations (ICC) = 28.7–40.5%, P < 0.0001), we used mixed regression modelling to nest births within each parent as random intercepts. The classical sandwich estimator (Diggle et al., 2002b) in variance estimation for fixed-effects was applied. Heterogeneity was tested using site-sex interaction term as a fixed effect. No statistical heterogeneity (P interaction > 0.05) was found for any of the models. We therefore carried out pooled analysis with adjustment for site and offspring sex as fixed effects.

We examined the associations of parental birthweight and childhood conditional linear growth with offspring LBW status. We used generalized linear mixed models (GLMM) to estimate the prevalence ratios (PR) of LBW adjusted for random effects with parental childhood conditional growth variables as the chief model predictors. Similarly, we related parental childhood stunting with offspring LBW status in a separate mixed logit model. Since prevalence estimates of LBW and parental childhood stunting were not low (13–84%), PR (95% CI) with robust variance (Zou, 2004) were considered to be more precise measures of association and were used instead of odds ratios.

**Sensitivity analyses**

All parental conditional variable derivation were based on birthweight rather than birthlength (unavailable for Brazil). The results of a sensitivity analysis showed very similar estimates using either birthlength or birthweight for the 3 cohorts that had both length and birthweight. Also, sensitivity analyses indicated no significant differences in results using growth measures obtained at 8.5 year (Philippines) vs. those obtained at 4 year (other 3 sites), as previously reported (Adair et al., 2013).

Offspring gestational age was only available for Guatemala and the Philippines and as a categorical variable (preterm: <37, term: 37–41 and post-term: >41 weeks) in India and was not included in the analysis. We conducted sensitivity analyses by first converting continuous gestational ages (GA) in Guatemala and the Philippines to categorical variables in order to carry out a pooled analysis with India. The results of this pooled analysis involving the three sites with GA were very similar to the main analyses with all four sites but not controlled for gestational age.

**Missing data**

About 22% of parents were missing one or more measure of childhood growth and SES. We used multiple imputation methods under an assumption of missingness at random and generated 15 datasets with imputed values (Little and Rubin, 2002) using the Amelia II package (Honaker et al., 2009). Imputation accuracy was assessed with over-impute diagnostics graph confidence bands and was found to be adequate (Honaker et al., 2009). Combined inference estimates were obtained from the 15 imputation datasets using Rubin’s Imputation inference rules. Sensitivity analyses conducted using mother-offspring or father-offspring pairs with available data (list-wise deletion) were identical to those obtained from imputation analyses, albeit with wider confidence intervals (results not shown). We therefore present results from the imputation analyses. Two-tailed P-values <0.05 were considered statistically significant. All analyses were conducted with PC-SAS version 9.3 (SAS Institute, Cary, NC), and R (R Core Team, 2014) software with the package ‘lme4’.

**RESULTS**

The ages of the parents at the last follow-up are shown in Table 1. Mean maternal and paternal adult heights were 156.6 cm and 164.4 cm, respectively. Prevalence of childhood stunting at 2 year ranged from 22 to 84% and was similar in fathers and mothers (Table 2). Low birthweight prevalence ranged between 11.5 and 20.0% across all sites for offspring of mothers and fathers. The mean number of births-to-date ranged from 1.5 for India to 2.2
in Guatemala. Fathers of the Brazilian and Indian cohorts reported fewer children in comparison to those from Guatemala and the Philippines.

Results of mixed linear models describing the association between parental growth in SD units and offspring birthweight are shown in Table 3. Also shown are the equivalents, in cm or g, corresponding to one SD unit of linear or relative weight, respectively. For mothers, 1 SD of conditional linear growth corresponded to 2.7 cm (birth–2 year), 3.5 cm (2–MC), and 5.4 cm (MC-adulthood) in absolute terms and for the fathers they were 4.0 cm (birth–2 year), 3.7 cm (2-MC) and 10.0 cm (MC-adulthood). Similarly, the absolute equivalents in grams per 1 SD of maternal birthweight was 478 g and that for her conditional relative weight gains were 847 g (birth–2 year), 1,072 g (2-MC), and 10,615 g (MC-adulthood).

Those of the fathers were 510 g/1SD of birthweight, and 921 g (birth–2 year), 1,182 g (2-MC), 10,524 g (MC-adulthood) per 1SD absolute weight gain. Absolute weight gains were higher (>10 kg) for the late childhood period (MC-Adulthood) and similar in both parents.

In adjusted models, maternal birthweight and postnatal conditional linear/length growth were associated with offspring birthweight. The increases per one SD were 102.3 g (95% CI: 79.5–125.2) with maternal birthweight, 45.5 g (95% CI: 20.8–70.3) with length growth 0–2 year, and 27.3 g (95% CI: 3.5–51.1) with height growth MC-Adulthood.

Paternal birthweight and conditional linear/length growth 0–2 year were associated with 57.3 g (95% CI: 25.9–88.6) and 55.5 g (95% CI: 20.8–90.3) increases in offspring birthweight, respectively; there were no associations with 2 year-MC or MC-adulthood. There were

## Table 2. Characteristics of parents and their offspring in four birth cohorts (3,392 parents; 5,506 offspring)

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Brazil mean (SD)/range</th>
<th>Guatemala mean (SD)/range</th>
<th>India mean (SD)/range</th>
<th>Philippines mean (SD)/range</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mothers (N = 2,031)</td>
<td>n = 1,063</td>
<td>n = 179</td>
<td>n = 285</td>
<td>n = 504</td>
</tr>
<tr>
<td>Firstborn status (%)</td>
<td>19.0</td>
<td>34.9</td>
<td>11.4</td>
<td>9.2</td>
</tr>
<tr>
<td>Birthweight (kg)</td>
<td>3.0 (0.5)</td>
<td>3.1 (0.5)</td>
<td>2.9 (0.4)</td>
<td>2.8 (0.5)</td>
</tr>
<tr>
<td>Stunted at 2 year (&lt;2SD) (%)</td>
<td>56.2</td>
<td>23.2</td>
<td>83.9</td>
<td>44.8</td>
</tr>
<tr>
<td>Adult height (cm)</td>
<td>156.6</td>
<td>160.0 (6.0)</td>
<td>150.5 (5.1)</td>
<td>155.3 (5.4)</td>
</tr>
<tr>
<td>Children per mother</td>
<td>1.5 (1–7)</td>
<td>1.6 (1–6)</td>
<td>2.2 (1–7)</td>
<td>1.5 (1–5)</td>
</tr>
<tr>
<td>Offspring of mothers N = 3,455</td>
<td>n = 1,942</td>
<td>n = 478</td>
<td>n = 918</td>
<td></td>
</tr>
<tr>
<td>Male (%)</td>
<td>54.0</td>
<td>51.1</td>
<td>53.7</td>
<td>56.0</td>
</tr>
<tr>
<td>Firstborn status (%)</td>
<td>52.7</td>
<td>61.0</td>
<td>37.5</td>
<td>58.2</td>
</tr>
<tr>
<td>Birth weight (kg)</td>
<td>3.0 (0.5)</td>
<td>3.2 (0.6)</td>
<td>3.0 (0.5)</td>
<td>2.8 (0.5)</td>
</tr>
<tr>
<td>Birth height &lt;2500 g (%)</td>
<td>16.1</td>
<td>11.5</td>
<td>15.8</td>
<td>17.0</td>
</tr>
<tr>
<td>Fathers N = 1,361</td>
<td>n = 596</td>
<td>n = 96</td>
<td>n = 342</td>
<td>n = 327</td>
</tr>
<tr>
<td>Firstborn status (%)</td>
<td>21.2</td>
<td>41.0</td>
<td>10.5</td>
<td>11.7</td>
</tr>
<tr>
<td>Birth weight (kg)</td>
<td>3.1 (0.5)</td>
<td>3.3 (0.5)</td>
<td>3.0 (0.4)</td>
<td>2.9 (0.4)</td>
</tr>
<tr>
<td>Stunted at 2 year (&lt;2SD) (%)</td>
<td>48.0</td>
<td>13.6</td>
<td>65.1</td>
<td>43.3</td>
</tr>
<tr>
<td>Adult height (cm)</td>
<td>164.4 (8.7)</td>
<td>173.5 (6.6)</td>
<td>151.3 (4.5)</td>
<td>169.9 (6.1)</td>
</tr>
<tr>
<td>Children per father</td>
<td>1.5 (1–7)</td>
<td>1.3 (1–4)</td>
<td>2.2 (1–7)</td>
<td>1.4 (1–4)</td>
</tr>
<tr>
<td>Offspring of fathers N = 2,051</td>
<td>n = 1,072</td>
<td>n = 564</td>
<td>n = 520</td>
<td></td>
</tr>
<tr>
<td>Male (%)</td>
<td>53.6</td>
<td>50.8</td>
<td>52.9</td>
<td>58.5</td>
</tr>
<tr>
<td>Firstborn status (%)</td>
<td>48.5</td>
<td>60.1</td>
<td>39.5</td>
<td>59.2</td>
</tr>
<tr>
<td>Birth weight (Kg)</td>
<td>3.0 (0.6)</td>
<td>3.3 (0.7)</td>
<td>2.9 (0.5)</td>
<td>2.8 (0.5)</td>
</tr>
<tr>
<td>Birth weight &lt;2500 g (%)</td>
<td>14.7</td>
<td>11.3</td>
<td>16.3</td>
<td>15.3</td>
</tr>
</tbody>
</table>

## Table 3. Associations between parental height and relative weight gain in specific periods of childhood and offspring birthweight in four birth cohorts (3,392 parents; 5,506 offspring)

<table>
<thead>
<tr>
<th>Parental growth measures</th>
<th>Equivalent per SD of conditional growth</th>
<th>Coefficient (95% CI)</th>
<th>P-value</th>
<th>Equivalent per SD of conditional growth</th>
<th>Coefficient (95% CI)</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Birthweight (z)</td>
<td>478 g</td>
<td>102.3 (79.5, 125.2)</td>
<td>&lt;0.01</td>
<td>510 g</td>
<td>57.3 (25.9, 88.6)</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td>Conditional height 0–2 year (z)</td>
<td>2.7 cm</td>
<td>45.5 (20.8, 70.3)</td>
<td>&lt;0.01</td>
<td>4.0 cm</td>
<td>55.5 (20.8, 90.3)</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td>Conditional height 2 year-MC (z)</td>
<td>3.5 cm</td>
<td>12.4 (10.6, 15.5)</td>
<td>0.29</td>
<td>3.7 cm</td>
<td>19.9 (13.7, 26.1)</td>
<td>0.16</td>
</tr>
<tr>
<td>Conditional height MC-adulthood (z)</td>
<td>5.4 cm</td>
<td>27.3 (5.5, 51.1)</td>
<td>0.03</td>
<td>10.0 cm</td>
<td>13.9 (17.6, 45.4)</td>
<td>0.39</td>
</tr>
<tr>
<td>Conditional relative weight gain 0–2 year (z)</td>
<td>847 g</td>
<td>16.6 (6.7, 39.8)</td>
<td>0.16</td>
<td>921 g</td>
<td>25.6 (4.7, 55.9)</td>
<td>0.10</td>
</tr>
<tr>
<td>Conditional relative weight gain 2 year-MC (z)</td>
<td>1,072 g</td>
<td>24.3 (1.5, 47.0)</td>
<td>0.04</td>
<td>1,182 g</td>
<td>−1.5 (−34.7, 31.6)</td>
<td>0.93</td>
</tr>
<tr>
<td>Conditional relative weight gain MC-adulthood (z)</td>
<td>10,615 g</td>
<td>49.1 (25.3, 72.8)</td>
<td>&lt;0.01</td>
<td>10,524 g</td>
<td>8.1 (−21.8, 38.1)</td>
<td>0.60</td>
</tr>
</tbody>
</table>

MC: Mid-childhood is 4 years for Brazil, Guatemala, and India and 8.5 yr for Philippines. Models are adjusted for parental early childhood SES (quintiles), maternal/ paternal firstborn status, offspring sex and firstborn status, nutrition supplementation status, site/4 categories), maternal age at delivery (<18, ≥18 year), and sibling clustering (with random intercepts). Estimates obtained from multiple imputation analyses of 15 replicate datasets.
significant associations between maternal relative weight for 2 year-MC ($\beta = 24.3 \text{ g}$, 95% CI: 1.3–47.0) and MC–adulthood ($\beta = 49.1 \text{ g}$, 95% CI: 25.3–72.8) and offspring birthweight. There were no significant associations with offspring birthweight for any of the paternal relative weight measures (Table 3).

Associations between parental conditional childhood growth and risk (prevalence ratio) of offspring LBW status are shown in Table 4. Parental birthweight and linear growth from 0 to 2 year were independently associated with reduced risk of offspring low birthweight, with prevalence ratio (PR) = 0.7 (95% CI: 0.6–0.8) and PR = 0.8 (95% CI: 0.7–0.9), respectively for mother-offspring; and PR = 0.87 (95% CI: 0.8–1.0) and PR = 0.80 (95% CI: 0.7–0.9), respectively for father-offspring models. Associations were nonsignificant for the middle and late periods of paternal linear growth. With respect to relative weight gain, there was a significant association with risk of LBW only for maternal conditional relative weight gain from 2 year–MC (PR = 0.9 (95% CI: 0.8–1.0)).

We also related parental childhood stunting to offspring birthweight and risk of LBW (Table 5). Maternal stunting at 2 year (compared to nonstunted) was associated with a 1.6-fold increase (PR = 1.6, 95% CI: 1.25–2.05) in the risk for offspring LBW and with a 107.6 g (95% CI: −160.3 to −54.8) decrease in offspring birthweight. Paternal childhood stunting was not significantly associated with offspring birthweight or LBW risk.

**DISCUSSION**

We find that parental birthweight and linear growth from birth to 2 year, but not later, are associated with offspring birthweight. Paternal relative weight gain, independent of linear growth, was not associated with offspring birthweight but maternal relative weight gain was significant for 2 year-MC and MC-adult, respectively. Finally, parental stunting at 2 year, a summary indicator of growth failure in early life, was associated with reductions in birthweight and was predictive of higher risk of low birthweight.

Associations between parental and offspring birthweight are well-known, with most such studies reporting a stronger relationship for maternal birthweight than for paternal birth weight (Kuzawa and Eisenberg, 2012; Mattsson and Rylander, 2013; Ramakrishnan et al., 1999). What is novel in our analysis is that linear growth from 0 to 2 year is strongly associated with offspring birthweight and that the strength of this relationship is similar for both mothers and fathers. Our results contradict a prior study from Brazil (Horta et al., 2009), which found a relationship between weight gain from 0 to 20 months in mothers but not fathers (Horta et al., 2009) (for firstborns only). In the present study, the specific findings for Brazil, with a much larger sample, 1,659 parents and 2,537 offspring (included all births and not firstborns only), are statistically significant for linear growth from 0 to 2 year for both fathers and mothers (results not shown).

The fact that the coefficients for relationships between linear growth (0–2 year, 2 year-MC and MC-adulthood; Table 3) and offspring birthweight are very similar for mothers and fathers indicates that, it is unlikely that the mechanisms operate solely through the maternal fetal environment. Since we use conditional variables that control for prior measures, we can exclude the influence of
likely to live in poor households as adults. Conversely, bet-
as compared to nonstunted adults. Women who were
spouses with shorter stature and lower grade attainment
stunted peers. Stunted adults married, on average,
scores for reading and intelligence as compared to non-
mala was associated with less schooling and lower test
in pooled models with categorized GA (for Guatemala,
Indian and Philippines) vs. all four sites pooled without
adjustment for GA. Even though we adjusted for several
confounding factors, residual confounding still remains a
possibility. Finally, since the parental cohorts married
mostly spouses from outside the cohorts, we had few cou-
ple who were both cohort members. This precluded many
types of analyses that would have been useful in teasing
apart paternal and maternal influences.

CONCLUSIONS

Parental birth weight and linear growth from birth to 2
years, but not later, and maternal conditional relative
weight (but not paternal) from 2 year through adulthood,
are associated with offspring birth weight These findings
suggest the importance of adequate nutrition during preg-
nancy and the first 2 year of life for both parents as well
as the role of preconception body composition in mothers.
Other mechanisms, such as shared genes, environment,
and epigenetic changes, may also contribute to associa-
tions between parental growth and offspring birthweight.

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butions: Addo, Martorell and Stein led the writing
team and planned the analyses. Addo, Martorell,
Osmond, and Stein analyzed the data. Addo, Martor-
ell, and Stein interpreted and wrote the first and last
manuscript drafts. Addo, Martorell, and Stein are
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ments on the manuscript and approved the final version.

The COHORTS group members are: Cesar G. Victora,
Fernando C. Barros, Denise Gigante, Pedro C. Hallal, and
Bernardo L. Horta (Universidade Federal de Pelotas, Bra-
zil); Reynaldo Martorell, Aryeh D. Stein, O.Yaw Addo,
Wei Hao (Hubert Department of Global Health, Rollins
School of Public Health, Emory University, Atlanta, GA,
USA); Manual Ramirez-Zea (Instituto de Nutricion de
Centra América y Panamá (INCAP), Guatemala City,
Guatemala); Santosh K. Bhargava (Sunderlal Jain Hospi-
tal, New Delhi, India); Harshpal Singh Sachdev, Shikha

The pattern of relationships with conditional relative
weight gain differs from that for linear growth. Father’s
conditional relative weight gain was not related with off-
spring birthweight and in mothers, coefficients are larger
for later compared to earlier changes. Of note is the
importance of relative weight gain from MC to adulthood,
which reflects acquisition of both lean and fat mass inde-
pendent of height and thus larger offspring birthweight (a
coefficient of 49 g/1 SD change). This potential benefit,
however, has to be balanced with the increased risks of
maternal obesity and chronic diseases that are associated
with relative weight at this period (Adair et al., 2013).
Our data came from low and middle income countries and
the degree to which they apply to industrialized societies,
which have much less growth failure, needs to be assessed
especially using conditional growth modeling for assess-
ning relative importance of parental childhood periods
of growth. Despite marked variations in early life growth
failure across the four sites, associations were similar
across sites; this suggests that our findings may indeed
apply to many settings.

The use of uncorrelated measures of growth for four
life course periods is a strength of this study that allowed
us to assess the relative importance of parental growth
during specific developmental periods. The use of condi-
tional height disentangled from weight gain allowed us to
isolate the role of linear growth from weight gain. The
inclusion of four cohorts that provide a sample size of 3,392
parents and 5,506 offspring afforded us statistical
power to test observed associations. Another strength of
this study is the use of mixed regression (mixed linear
and logit) models that controlled for any potential intra
family correlations in birthweight (Diggle et al., 2002a;
Fritzmaurice et al., 2004).

Several limitations of this study warrant discussion.
One limitation of our study is that offspring birthweight
was obtained largely by recall. This may have contributed
to loss of precision and underestimation of the associa-
tions. Birth length data were unavailable for Brazil and
as a result we used conditional growth measures condi-
tioned on birthweight and all prior heights for all the four
study sites instead. Gestational age (GA) was unavailable
for Brazil. Sensitivity analyses indicated similar results
in pooled models with categorized GA (for Guatemala,
India and Philippines) vs. all four sites pooled without
adjustment for GA. Even though we adjusted for several
confounding factors, residual confounding still remains a
possibility. Finally, since the parental cohorts married
mostly spouses from outside the cohorts, we had few cou-
ple who were both cohort members. This precluded many
types of analyses that would have been useful in teasing
apart paternal and maternal influences.
Sinha (Sitaram Bhartia Institute of Science and Research, New Delhi, India); Poornima Prabhakaran (Public Health Foundation of India, New Delhi, India); Clive Osmond, Caroline Fall/MRC LifeCourse Epidemiology Unit, University of Southampton, Southampton, UK), Linda Adair (Department of Nutrition, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA); Isabelita Bas, Nanette Lee, Judith Borja (Office of Population Studies, University of San Carlos, Cebu City, Philippines); Darren Dahluy (University of Leeds, Leeds, U.K.); Christopher Kuzawa (Department of Anthropology, Northwestern University, Evanston, IL, USA); Linda Richter, Shane Norris, Julia De Kadt (Developmental Pathways for Health Research Unit, Department of Paediatrics, University of the Witwatersrand, Johannesburg, and South Africa), Alan Stein (Department of Psychiatry, University of Oxford).

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**LITERATURE CITED**


